

# Rain Gauge Anomaly Detection Using Gaussian Processes

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## 1 Introduction

We aim to explore the use of Gaussian processes to model rainfall potential and exploit this model to detect anomalous rain gauge readings. Faulty readings commonly appear as:

1. Null readings when rainfall occurred
2. Extreme readings when lower rainfall occurred

Let us focus on the initial case, zero readings when the rainfall was greater than zero.

### 1.1 Model

The rainfall is observed at  $L$  locations,  $\mathbf{x} = \{\mathbf{x}_1, \dots, \mathbf{x}_L\}$ , over  $T$  time intervals with the potential for rain at time  $t$

$$f_t(\mathbf{x}) \sim \mathcal{GP}(\mathbf{0}, \mathbf{K}_t). \quad (1)$$

The rainfall  $y_{it}$  observed at location  $i$  at time  $t$  is given as follows

$$y_{it} = \begin{cases} f_{it} + \epsilon_{it} & \text{w.p. } \theta_i \\ \epsilon_{it} & \text{w.p. } 1 - \theta_i \end{cases} \quad (2)$$

where  $\epsilon_{it} \sim \mathcal{N}(0, \sigma_\epsilon^2)$  is observation noise. If we introduce a Bernoulli random variable  $d_{it} \sim \mathcal{B}(\theta_i)$ , we can write the observation model as

$$y_{it} = f_{it}d_{it} + \epsilon_{it} \quad (3)$$

Let  $\mathbf{y}_t = [y_{1t}, \dots, y_{Lt}]^T$  and similarly introduce  $\mathbf{f}_t, \mathbf{d}_t, \boldsymbol{\theta}$ .

The joint distribution for our problem is given by

$$p(\mathbf{Y}, \mathbf{F}, \mathbf{D} | \boldsymbol{\theta}, \gamma) = \prod_{t=1}^T p(\mathbf{y}_t, \mathbf{f}_t, \mathbf{d}_t | \boldsymbol{\theta}, \gamma) \quad (4)$$

Where  $\gamma$  are the hyper-parameters controlling the nature of  $\mathbf{K}_t$ .  
 Let's assume  $T = 1$ . We have

$$p(\mathbf{y}_t, \mathbf{f}_t, \mathbf{d}_t | \boldsymbol{\theta}, \boldsymbol{\gamma}) = p(\mathbf{y}_t | \mathbf{f}_t, \mathbf{d}_t) p(\mathbf{f}_t | \boldsymbol{\gamma}) p(\mathbf{d}_t | \boldsymbol{\theta}) \quad (5)$$

where

$$p(\mathbf{y}_t | \mathbf{f}_t, \mathbf{d}_t) = \prod_{i=1}^L \mathcal{N}(y_{it} | d_{it} f_{it}, \sigma_\epsilon^2),$$

$$p(\mathbf{d}_t | \boldsymbol{\theta}) = \prod_{i=1}^L \theta_i^{d_{it}} (1 - \theta_i)^{1-d_{it}},$$

and

$$p(\mathbf{f}_t | \boldsymbol{\gamma}) = \mathcal{N}(\mathbf{f}_t | \mathbf{0}, \mathbf{K}_t).$$

## 1.2 Inference

We would like to determine:

- The parameters  $\boldsymbol{\gamma}, \boldsymbol{\theta}$
- The posterior  $p(\mathbf{f}_t | \mathbf{y}_t)$
- The predictive distribution  $p(\mathbf{f}_t^* | \mathbf{y}_t)$

To get the parameters, we would need to maximize the marginal likelihood

$$\boldsymbol{\gamma}^*, \boldsymbol{\theta}^* = \arg \max_{\boldsymbol{\theta}, \boldsymbol{\gamma}} \sum_{\mathbf{d}_t} \int p(\mathbf{y}_t, \mathbf{f}_t, \mathbf{d}_t | \boldsymbol{\theta}, \boldsymbol{\gamma}) d\mathbf{f}_t$$

This is intractable due to the sum over all possible  $\mathbf{d}_t$ . We proceed via variational inference where we optimize the lower bound on the marginal likelihood.

$$\log p(\mathbf{y}_t | \boldsymbol{\theta}, \boldsymbol{\gamma}) \geq \mathbb{E}\{\log p(\mathbf{y}_t, \mathbf{f}_t, \mathbf{d}_t | \boldsymbol{\theta}, \boldsymbol{\gamma})\} - \mathbb{E}\{\log q(\mathbf{f}_t, \mathbf{d}_t)\} \quad (6)$$

where

$$q(\mathbf{f}_t, \mathbf{d}_t) = q(\mathbf{f}_t) \prod_{i=1}^L q(d_{it})$$

is the variational approximation to the true posterior.

We proceed as in [1] to derive the variational approximation to the posterior as well as the hyperparameters. We find that

$$q(\mathbf{f}_t) = \mathcal{N}(\mathbf{f}_t | \boldsymbol{\mu}_f, \boldsymbol{\Sigma}_f)$$

and

$$q(d_{it}) = \eta_i^{d_{it}} (1 - \eta_i)^{1-d_{it}}$$

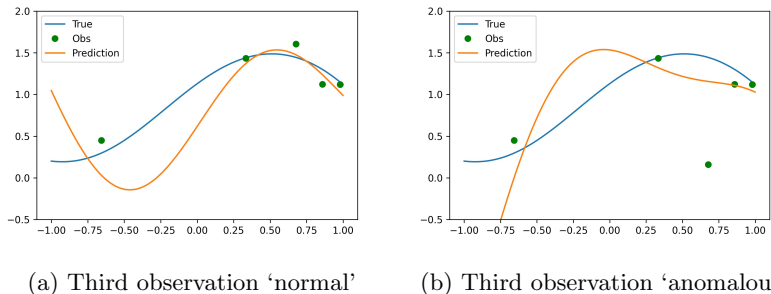


Figure 1: Latent process observed through noise with probability of being switched off.

### 1.3 Experiments

We explore  $\theta^*$  and the predictive mean in a 1D example where the observations are drawn from a GP with a squared exponential kernel. See Figure 1. We compare the values of  $\theta^*$  in a case with an anomaly and without. For the observations in Figure 1a where there are no anomalies,  $\theta^* = [0, 1, 1, 1, 1]^T$  while in Figure 1a,  $\theta^* = [0, 1, 0, 1, 1]^T$  where the third observation is anomalous.

#### 1.3.1 Observations

We note that:

1. The posterior mean appears to ‘ignore’ the anomalous observation which is promising
2. ‘Low’ values are assumed anomalous. It appears that the noise variance is over estimated leading to a lack of identifiability.

#### 1.3.2 Next steps

1. Investigate the estimation of noise variance
2. Verify VB implementation by monitoring the evidence lower bound (ELBO)
3. Experiments with real data

## References

- [1] Michalis Titsias and Miguel Lázaro-Gredilla. Spike and slab variational inference for multi-task and multiple kernel learning. *Advances in neural information processing systems*, 24, 2011.